

Genetic and Environmental Influences on Adult Mental Health: Evidence for Gene-Environment Interplay as a Function of Maternal and Paternal Discipline and Affection

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Abstract Researchers have long theorized that genetic influence on mental health may differ as a function of environmental risk factors. One likely moderator of genetic and environmental influences on psychopathological symptoms is parenting behavior, as phenotypic research shows that negative aspects of parent–child relationships are associated with greater likelihood of mental illness in adulthood. The current study examined whether levels of reported parental discipline and affection experienced in childhood act as a trigger, or buffer, for adult mental health problems. Results from a nationwide twin sample suggest level of father’s discipline and affection, as reported by now-adult twins, moderated genetic and environmental influences on internalizing symptoms in adulthood, such that heritability was greatest at the highest levels of discipline and affection. Father’s affection also moderated the etiological influences on alcohol use problems, with greater heritability at the lowest levels of affection. No moderating effect was found for mothers. Findings suggest relationships with fathers in childhood can have long-lasting effects on the etiological influences on adult mental health outcomes.

Keywords Mental health · Parent–child relationships · Gene-environment interplay · Twin · Behavior genetics

All common forms of psychopathology have a significant heritable component (Plomin et al. 2012), and certain risk or

protective factors may moderate genetic influences on mental health outcomes (e.g., Cadoret et al. 1995). For more than 10 years, researchers have utilized twin data to examine whether heritability estimates for different types of psychopathology vary as a function of “moderator” variables (Purcell 2002). Consistent with this idea, socioeconomic status (SES) moderates genetic influences on antisocial behavior (Tuvblad et al. 2006) and internalizing psychopathology (South and Krueger 2011). In both cases, greater genetic influences are found at the highest ends of SES, evidence of what Bronfenbrenner and Ceci (1994) referred to as a bioecological model or what others (Raine 2002) have called a *social push* theory, suggesting that heritability is greatest in the most enriched environment and lowest in the most disadvantaged environment. In other work, researchers have found that genetic influences are highest in the most distressed environments (South and Krueger 2008), evidence of a diathesis-stress model in which latent risk for psychopathology is triggered by stressful events (Monroe and Simons 1991). Finally, it is also possible for biometric moderation models to show that genetic influences are elevated at the extreme levels of the moderator (South and Krueger 2013), evidence of differential susceptibility (Belsky and Pluess 2009) or the “orchid” model (Ellis and Boyce 2008), the idea being that some individuals, like dandelions, will flourish anywhere, while some, like an orchid, need certain conditions to thrive. The *quantitative* models that test gene X environment interaction have the advantage of also modeling changes in environmental influences on psychopathology; going back to the example of SES and mental health, the greatest influence of the unique environment on internalizing psychopathology was at the extreme low end of SES (South and Krueger 2011).

One likely moderator of genetic and environmental influences on symptoms of psychopathology is the parent–

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child relationship, because parenting styles can place children at risk for psychopathology. In particular, parenting that is rejecting or overcontrolling has been linked to later mental illness (e.g., Parker et al. 1995). Retrospective reports of authoritarian parenting have been linked to later well-being and depressive symptoms but not substance abuse (Rothrauff et al. 2009). In addition to having phenotypic, correlational relationships with psychopathology, this type of parenting may work on a different level, moderating the etiological influences on mental illness.

A handful of studies have examined parent–child relationships as a moderator of the genetic and environmental influences on mental health outcomes. For example, Hicks et al. (2009b) reported that genetic variance of externalizing psychopathology was greater for adolescents who reported the most problems in relationships with their parents. This would support a diathesis-stress model of externalizing problems, with environmental adversity triggering expression of genetic influences. However, the findings across studies are inconsistent, possibly because researchers have not distinguished, first, the impact of relationships with mothers and fathers (e.g., Feinberg et al. 2007), and second, the two different dimensions of discipline and affection (e.g., Hicks et al. 2009a, b).

Even when separate studies use putatively similar parenting variables, results can be conflicting. In two separate studies, maternal discipline had opposite effects on internalizing and externalizing behavior. One study found the greatest genetic influences on internalizing problems (i.e., depressive symptoms) at the highest levels of punitive discipline (Lau and Eley 2008). This would also support a diathesis-stress model of psychopathology if the high punitive discipline environment is interpreted as the more “at-risk” environment. Another study using the same sample, however, found greater influence of genetics in the low maternal punitive discipline environment for externalizing behavior (Button et al. 2008), supportive of the social-push theory of psychopathology. Still other work has reported that maternal and paternal discipline had opposite effects on the genetic variance in externalizing behavior (Button et al. 2008), with greater genetic variance found at low levels of maternal discipline and high levels of paternal discipline.

The conflicting findings in the literature to date may result from a variety of method-related factors, including different ages of twin participants, assessment of psychopathology, and even the type of moderation model used. Indeed, recent work on biometric moderation suggests that failure to consider nonlinear effects may result in spurious evidence of G \times E (Van Hulle et al. 2013). It is possible, however, that parenting may be a significant biometric moderator of psychopathology, and the mixed findings from studies conducted so far may reflect that

parent–child relationships moderate genetic influence at the level of specific combinations of parent (mother, father), dimension of parental behavior (affection, discipline), and psychopathology (internalizing symptoms, alcohol use problems). These levels have not been examined in previous work. This is a key limitation that prevents us from understanding when and under what circumstances the parent–child relationship can serve as a trigger or a buffer for the expression of genetic and environmental influences on psychopathology.

The aim of the current research was to examine components of the parent–child relationship that moderate genetic and environmental influence on adult symptoms of psychopathology. Previous research has exclusively examined childhood and adolescent samples, leaving unanswered the question of whether parenting can have a lasting impact on the etiology of mental health into adulthood. The current study used an exploratory approach to try and untangle mixed findings from previous work. Retrospective reports of parental discipline and affection were obtained, separately for mothers and fathers, from monozygotic (MZ) and dizygotic (DZ) twins. Each combination of parent and parental dimension (e.g., mother’s discipline) was examined as a potential moderator for the genetic and environmental influences on an internalizing factor score and a sum score of alcohol use problems.

Given the conflicting results in the literature, and the exploratory nature of the current study, predictions about the direction of effects were not hypothesized but it was expected that moderation would be found for some, but not all, combinations of mental health problems, parent, and dimension of parenting. If genetic influences on symptoms of psychopathology were greatest in the most “enriched” environment (i.e., one marked by high affection and low levels of strict discipline) a social push theory of psychopathology would be supported, whereas if genetic influences were greatest in the most “risky” environment (i.e., with lowest levels of affection and strictest discipline) it would be evidence in favor of a diathesis-stress model. It is also possible that genetic influences might be elevated at extreme ends of the moderator variable, evidence of a *differential susceptibility/orchid model*. If differential susceptibility was found for discipline, for instance, it might suggest that either too little or too much discipline leads to a greater expression of genetic influences.

Method

Sample

Participants were drawn from a nationwide twin sample recruited for the MacArthur Foundation Survey of Midlife

Development in the United States (MIDUS). The larger MIDUS sample is a nationally representative cohort of individuals aged 25–74 years drawn from the non-institutionalized civilian population of the continental United States. The MIDUS study twin subsample (Kessler et al. 2004) was ascertained through a telephone survey of approximately 50,000 households (Kessler et al. 2004). The final response rate (i.e., both members of the twin pair were contacted by an interviewer, agreed to participate, and completed a short zygosity screening questionnaire) was 26 %. The twin subsample completed the same MIDUS battery as the full sample, consisting of a computer-assisted telephone interview and two mailed questionnaire booklets (Kendler et al. 2000).

A total of 1996 people were recruited for and enrolled in the MIDUS twin sample. Twins completed a brief twin screen to determine zygosity, a technique that is generally more than 90 % accurate (cf. Lykken et al. 1990). Twins with missing or indeterminate zygosity information, twins who only completed the twin screener and no other measures, and opposite-sex twins were eliminated. This left a total of 1386 twins. The biometric models used here require data from both twins on the moderator variable; some twins were missing all data for a paired co-twin, leaving a total of 1344 twins from 672 twin pairs used in the biometric models: 164 monozygotic (MZ) male pairs, 186 MZ female pairs, 124 dizygotic (DZ) male pairs, and 198 DZ female pairs. The average age of the sample of complete twin pairs was 45 ($SD = 12.17$, range = 25–74).

Measures

Parent–child relationship quality The MIDUS questionnaire obtained retrospective reports of parental affection and discipline received during childhood. Seven items each retrospectively assessed maternal and paternal affection. One of the items (*How would you rate your relationship with your mother/father during the years you were growing up?*) was rated on a 5-point scale ($1 = Excellent$, $5 = Poor$). The other six items (i.e., *How much did s/he understand your problems and worries?*, *How much could you confide in him/her about things that were bothering you?*, *How much love and affection did s/he give you?*, *How much time and attention did s/he give you when you needed it?*, *How much effort did s/he put into watching over you and making sure you had a good upbringing?*, *How much did s/he teach you about life?*) were rated on a 4-point scale ($1 = A lot$, $4 = Not at all$). A total scale score was calculated by taking the mean of the 7 items (the first item was multiplied by .75 to maintain continuity with the other six variables). Four items were included to assess maternal and paternal discipline by examining how strict, consistent, and harsh parents were during childhood (Rossi

2001). These items (i.e., *How strict was s/he with her rules for you?*, *How consistent was s/he about the rules?*, *How harsh was s/he when s/he punished you?*, *How much did s/he stop you from doing things that other kids your age were allowed to do?*) were rated on a 4-point scale ($1 = A lot$, $4 = Not at all$), and total scores were calculated as means of the 4 items.

There was missing data for reports of maternal and paternal discipline/affection (maternal discipline = 7.5 % missing, maternal affection = 7.4 % missing, paternal discipline = 12.4 % missing, paternal affection = 12.0 % missing). Because the biometric moderation model used in the current analyses requires that both twins have data on the moderator variable, mean substitution techniques were used to account for the missing data and maintain the sample size. Higher scores indicate more discipline and more affection, respectively. The average score for maternal discipline was 2.95 ($SD = .61$, range = 1–4, $\alpha = .76$). The mean for maternal affection was 3.17 ($SD = .66$, range = .96–3.96, $\alpha = .91$). Paternal discipline and affection had average scores of 2.98 ($SD = .71$, range = 1–4, $\alpha = .84$) and 2.73 ($SD = .74$, range = .75–3.96, $\alpha = .92$), respectively.¹

Internalizing symptoms Internalizing scores were taken from three DSM-III-R disorders, major depressive episode (MD), generalized anxiety disorder (GAD), and panic disorder (PD), as well as a rating of neuroticism. The 12-month prevalence for symptoms of the three disorders were measured during the phone interview using Composite International Diagnostic Interview Short Form scales (CIDI-SF; Kessler et al. 1998b). Research demonstrates good total classification accuracy (percentage of respondents whose CIDI-SF classification is the same as their classification of the full CIDI; Kessler et al. 1998a) and agreement with clinical diagnoses (Kessler et al. 1998b; Wittchen 1994). Neuroticism was assessed on the MIDUS self-administered questionnaire, using a personality measure based on the Five Factor Model (FFM; Lachman and Weaver 1997). The Neuroticism scale includes four adjectives rated on a 1–4 scale ($1 = A lot$, $4 = Not at all$): ‘moody,’ ‘worrying,’ ‘nervous,’ and ‘calm’ (reverse scored). Items were averaged to create a total neuroticism score.

As detailed elsewhere (see South and Krueger 2011), symptom counts for GAD, MD, and PD, and the scale score for neuroticism, were entered into a one-factor

¹ Means were taken from the sample of 1386 participants, who had known zygosity, information from multiple measures, and a same-sex co-twin. These means were used in the mean substitution procedures. Analyses were also conducted using means from the sample of 1344 twins, where each twin had data for a co-twin for the mean substitution procedures. The latter did not influence the findings of the study.

confirmatory factor analysis conducted in Mplus (Muthén and Muthén 1998–2012). Factor scores were extracted for use in the biometric moderation analyses and reversed to be positively correlated with the parenting variables (e.g., higher parental affection was positively correlated with less internalizing).

Alcohol use problems Alcohol use problems were assessed for the previous 12 months using the Alcohol Screening Test (AST; Selzer 1971) and two other questions (i.e., “using larger amounts longer than intended,” and “suffering the effects or after-effects of alcohol at home or work”). The AST consists of five dichotomous (*Yes/No*) items asking if alcohol use heightened chances of getting hurt, if emotional or psychological problems resulted from alcohol use, if there was a strong desire or urge to use alcohol, if a great deal of time was spent using or recovering from the effects of alcohol, and if more alcohol than usual had to be used to get the same effect. The additional alcohol use questions were rated on a 6-point scale (1 = *never*, 2 = *once or twice*, 3 = *3 to 5 times*, 4 = *6 to 10 times*, 5 = *11 to 20 times*, 6 = *more than 20 times*) but dichotomized (*Never/Occurred*) and summed with the other items (see Jarnecke and South 2014). Raw scores for alcohol use problems ranged from 0 to 7 and participants were required to respond to at least five of the seven items examined for analysis. Higher scores reflected more alcohol use problems. The average score for alcohol problems was .50 ($SD = 1.08$, range = 0–7, $\alpha = .73$). Scores were inversely transformed (and thus reversed) for use in the biometric analyses to adjust for negative skew.

Data analysis

The current analyses utilized biometric models designed for twin data. In general, biometric modeling with twin data uses the differences in MZ twins, who share 100 % of their genes, and DZ twins, who share 50 % of their independently segregating genes, to decompose the variance in a phenotype, like mental illness, into the variance due to genetic effects (A), common environmental influences (C), and unique environmental influences (E).

To determine the degree to which the ACE estimates on internalizing symptoms and alcohol use problems were shared with those contributing to parenting behavior, bivariate (Cholesky) decompositions were tested first. This bivariate decomposition includes latent factors representing the additive genetic, shared environmental, and nonshared environmental variance in the phenotype of interest (here, internalizing or alcohol problems) as well as latent factors representing the variance shared between the parenting behavior variable and the outcome psychopathology variable. To allow for the possibility that ACE estimates of internalizing symptoms and alcohol use problems vary as a function

of parental behavior (abbreviated here for simplicity as GxM), moderation terms were included in the second set of models (Van Hulle et al. 2013).² For each of the paths influencing the outcome variable, this bivariate Cholesky with GxM includes an overall coefficient separate from the moderator variable (i.e., parenting behavior) that indicates the magnitude of each effect of A, C, or E on the outcome, and the product of a coefficient that indexes the moderation of the outcome by parenting behavior multiplied by the level of the moderator (see Fig. 1a). Parameter estimates from the model are then used to plot the model-predicted genetic and environmental components of variance for psychopathological symptoms at different levels of parental behavior.

Because the biometric moderation model described above may be prone to false positives (van der Sluis et al. 2012), several alternative models were tested as well. The first of these alternative models was a nonlinear main effects model with GxM (Van Hulle et al. 2013; see Fig. 1b). This model removes moderation paths common to the moderator and outcome variables but retains the unique moderation paths, while including the linear and nonlinear main effects (β_1 and β_2) of the moderator on the outcome. This allows for the linear and nonlinear effect of the moderator on the phenotype to be modeled directly rather than through common moderation paths. Next, a more restrictive, nonlinear main effects model was fit, which constrained the unique moderation paths to zero. Finally, a linear main effects model was tested that examined the direct effect of the moderator on the outcome variable, allowing for the ACE decomposition of both the moderator and outcome variables.

Parenting behavior and psychopathology (i.e., internalizing symptoms and alcohol use problems) variables were regressed on age, age^2 , $age \times gender$, and $age^2 \times gender$ (McGue and Bouchard 1984) to correct for potential biases in model fitting due to gender and age differences. The residuals were then used to fit models in Mplus (Muthén and Muthén 1998–2012) using full-information maximum-likelihood to account for missing data in the outcome variable. Model fit was evaluated using the Akaike Information Criterion (AIC; Akaike 1987), Bayesian Information Criterion (BIC), and the likelihood ratio test (LRT). The difference in the -2 log-likelihood values for two separate models, the LRT is distributed as Chi square; a statistically significant difference in LRT between two models can indicate improvement in the model's fit as a result of adding or removing parameters. The AIC and BIC are information theoretic fit statistics that balances fit and number of parameters when choosing the best-

² Biometric moderation was also tested using alternative modeling procedures (see Purcell 2002) in Mx. Results were largely consistent with what is reported here, in terms of model fit and variance components. Full results are available from the first author. Note that the GxM model tested here is equivalent to the bivariate biometric moderation model described by Purcell (2002).

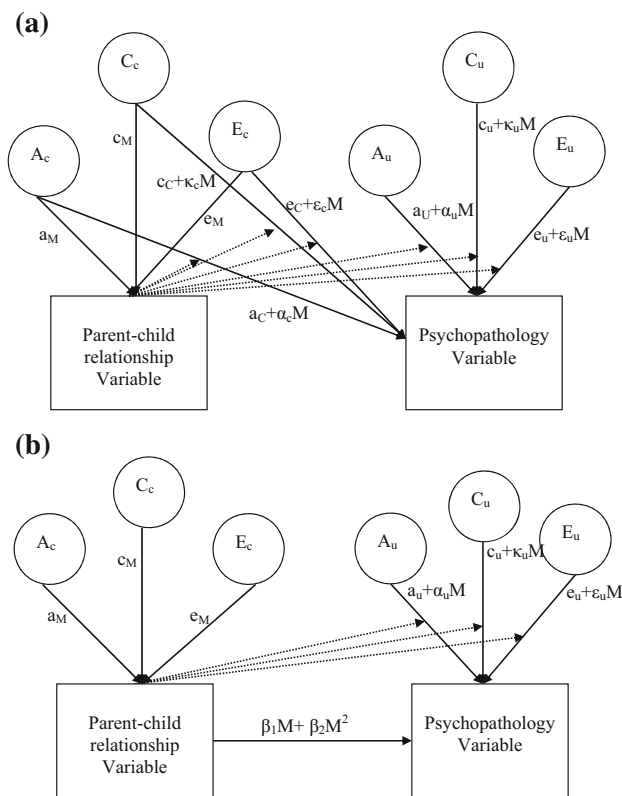


Fig. 1 **a** Bivariate Cholesky with GxM model with the parent–child relationship variable moderating genetic and environmental effects on the mental health variable (model displayed for single member of a twin pair). A additive genetic effects, C shared environmental variance, E nonshared environmental influences. a_M , c_M , and e_M represent genetic, shared environmental, and nonshared environmental parameter estimates, respectively, for the moderator variable. a_c , c_c , and e_c signify variance shared between the moderator variable and the outcome variable. a_u , c_u , and e_u represent residual variance unique to the outcome variable. **b** Nonlinear main effects with GxM model with the parent–child relationship variable moderating genetic and environmental effects on the mental health variable (displayed for one member of a twin pair). A additive genetic effects, C shared environmental variance, E nonshared environmental influences. a_M , c_M , and e_M represent genetic, shared environmental, and nonshared environmental estimates for the moderator variable. a_u , c_u , and e_u denote residual variance unique to the outcome variable. α_u , κ_u , and ε_u denote interactions between the moderator (M) and the genetic and environmental influences on the outcome variable. β_1 and β_2 represent the linear and non-linear main effects of the moderator variable on the outcome variable

fitting model, with lower values indicating better fitting models (Markon and Krueger 2004).

Results

Parenting behavior and internalizing symptoms

The raw internalizing factor score (not reversed) was significantly correlated with raw scores for maternal affection

($r = -.18$, $p < .001$) but not maternal discipline ($r = .00$, $p = 1.00$). The internalizing score was correlated with paternal affection ($r = -.15$, $p < .001$) but not paternal discipline ($r = .03$, $p = .25$). The five models described above (and listed in Table 1) were fit to the data for each combination of parenting variable and internalizing: bivariate Cholesky (BivCholesky), bivariate Cholesky with GxM (BivCholesky GxM), nonlinear main effects model with GxM (NL Main GxM), nonlinear main effects model (NL Main), and linear main effects model (Linear Main).

When maternal discipline and internalizing were entered into the models, the NL Main model provided the best fit to data according to AIC (7397) and BIC (7442; see Table 1). The NL Main model resulted in a nonsignificant loss of fit from the NL Main GxM model ($\chi^2_{diff} = 2.85$, 3 df, $p = .42$), but moving to the Linear Main model significantly worsened model fit ($\chi^2_{diff} = 124.27$, 1 df, $p = .00$). When the models were fit to the maternal affection and internalizing variables, the NL Main model again provided the best fit according to LRT and BIC ($\chi^2_{diff} = 5.04$, 3 df, $p = .17$; BIC = 7299). The AIC was lowest for the BivCholesky GxM (7252) but close to the NL Main model (7254). These results indicated that the genetic and environmental components of internalizing symptoms did not vary by level of perceived maternal discipline and affection, after accounting for the linear and nonlinear effects of maternal variables on internalizing.

Because there was no support for moderation, the ACE estimates for these models were interpreted at the average levels of maternal discipline and affection, respectively (see Table 2). For the maternal discipline NL Main model, genetic influences accounted for 21 % of the variance in internalizing, shared environmental influences accounted for 14 %, and nonshared environmental influences accounted for 64 % of the variance. Forty-four percent of the variance in maternal discipline was due to genetic effects, 11 % was due to shared environmental effects, and 45 % was due to nonshared environmental effects. Estimates for internalizing from the maternal affection model suggested that 14, 18, and 64 % of the variance in internalizing was accounted for by genetic, shared environmental, and nonshared environmental influences, respectively. Forty-four, 23, and 34 % of the variance in maternal affection was attributed to genetic, shared environmental, and nonshared environmental influences, respectively.

Next, the models were fit to paternal variables and internalizing. For paternal discipline and internalizing, the NL Main GxM provided the best fit to data according to AIC (7360) and BIC (7418). Moving from the NL Main GxM model to the NL Main model also resulted in a significant decrement in fit ($\chi^2_{diff} = 15.17$, 3 df, $p = .00$). This suggests that paternal discipline moderated the genetic and

Table 1 Fit statistics for biometric moderation models

	$-2\ln L$	Model comparison	$\Delta\chi^2$	Δdf	p	AIC	BIC
<i>Internalizing</i>							
Maternal discipline							
1. BivCholesky	-3812.70					7647	7698
2. BivCholesky GxM	-3685.01	Model 2 versus 1	127.69	6	0.00	7406	7487
3. NLMain GxM	-3685.71	Model 3 versus 2	0.70	4	0.95	7397	7456
4. NL Main	-3688.55	Model 4 versus 3	2.85	3	0.42	7397	7442
5. Linear Main	-3812.82	Model 5 versus 4	124.27	1	0.00	7644	7685
Maternal affection							
1. BivCholesky	-3736.41					7495	7545
2. BivCholesky GxM	-3607.85	Model 2 versus 1	128.57	6	0.00	7252	7333
3. NLMain GxM	-3611.88	Model 3 versus 2	4.03	4	0.40	7250	7308
4. NL Main	-3616.92	Model 4 versus 3	5.04	3	0.17	7254	7299
5. Linear Main	-3738.77	Model 5 versus 4	121.85	1	0.00	7496	7537
Paternal discipline							
1. BivCholesky	-3807.56					7637	7687
2. BivCholesky GxM	-3663.99	Model 2 versus 1	143.57	6	0.00	7364	7445
3. NLMain GxM	-3666.90	Model 3 versus 2	2.91	4	0.57	7360	7418
4. NL Main	-3682.07	Model 4 versus 3	15.17	3	0.00	7384	7429
5. Linear Main	-3808.85	Model 5 versus 4	126.78	1	0.00	7636	7677
Paternal affection							
1. BivCholesky	-3722.34					7467	7517
2. BivCholesky GxM	-3591.48	Model 2 versus 1	130.86	6	0.00	7219	7300
3. NLMain GxM	-3594.54	Model 3 versus 2	3.06	4	0.55	7215	7274
4. NL Main	-3599.99	Model 4 versus 3	5.45	3	0.14	7220	7265
5. Linear Main	-3723.46	Model 5 versus 4	123.47	1	0.00	7465	7506
<i>Alcohol use problems</i>							
Maternal discipline							
1. BivCholesky	-3659.74					7341	7392
2. BivCholesky GxM	-3551.29	Model 2 versus 1	108.45	6	0.00	7139	7220
3. NLMain GxM	-3551.96	Model 3 versus 2	0.67	4	0.96	7130	7189
4. NL Main	-3552.51	Model 4 versus 3	0.55	3	0.91	7125	7170
5. Linear Main	-3660.24	Model 5 versus 4	107.73	1	0.00	7339	7380
Maternal affection							
1. BivCholesky	-3602.02					7226	7276
2. BivCholesky GxM	-3491.12	Model 2 versus 1	110.90	6	0.00	7018	7099
3. NLMain GxM	-3491.77	Model 3 versus 2	0.65	4	0.40	7010	7068
4. NL Main	-3494.03	Model 4 versus 3	2.25	3	0.17	7008	7053
5. Linear Main	-3602.21	Model 5 versus 4	108.19	1	0.00	7222	7264
Paternal discipline							
1. BivCholesky	-3653.88					7330	7380
2. BivCholesky GxM	-3541.28	Model 2 versus 1	112.60	6	0.00	7119	7200
3. NLMain GxM	-3544.40	Model 3 versus 2	3.12	4	0.54	7115	7173
4. NL Main	-3544.87	Model 4 versus 3	0.47	3	0.92	7110	7155
5. Linear Main	-3656.57	Model 5 versus 4	111.71	1	0.00	7331	7372
Paternal affection							
1. BivCholesky	-3578.87					7180	7230
2. BivCholesky GxM	-3460.87	Model 2 versus 1	118.00	6	0.00	6958	7039
3. NLMain GxM	-3461.90	Model 3 versus 2	1.03	4	0.90	6950	7008

Table 1 continued

	$-2\ln L$	Model comparison	$\Delta\chi^2$	Δdf	p	AIC	BIC
4. NL Main	–3467.60	Model 4 versus 3	5.70	3	0.13	6955	7000
5. Linear Main	–3579.65	Model 5 versus 4	112.05	1	0.00	7177	7218

Best fitting model shown in bold. *BivCholesky* bivariate Cholesky, *BivCholesky GxM* bivariate Cholesky with GxM, *NL main GxM* nonlinear main effects model with GxM, *NL Main* nonlinear main effects model, *linear main* linear main effects model, $-2\ln L$ -2 loglikelihood, df degrees of freedom, *AIC* Akaike information criterion, *BIC* Bayesian information criterion

environmental influences unique to internalizing symptoms. When the parameter estimates from the NL Main GxM model were plotted, heritability estimates increased from low ($a^2 = 2\%$) to high ($a^2 = 48\%$) levels of paternal discipline (see Table 2; Fig. 2a). Influences from the shared environment decreased from 53% at low levels of discipline to 1% at +1 SD of discipline before increasing to 8% at high levels (+2SD) of discipline. Finally, the proportion of variance from the unique environment was curvilinear, increasing from low ($e^2 = 45\%$) to high ($e^2 = 66\%$) levels of paternal discipline before decreasing at higher levels of paternal discipline ($e^2 = 44\%$). Confidence intervals around the moderation parameter estimates suggest that much of the moderation seems to be occurring on the shared environmental path (see Table 3). This model also provided ACE estimates for paternal discipline. The model showed that 19% of the variance in paternal discipline was explained by genetic influences, 31% by shared environmental influences, and 48% by nonshared environmental influences.

For paternal affection and internalizing, the NL Main GxM model provided the best fit to the data according to AIC (7215). However, moving to the NL Main model resulted in a nonsignificant loss of fit ($\chi^2_{diff} = 5.45$, 3 df, $p = .14$). BIC was better for the NL Main model (7265) than the NL Main GxM model (7274) but by less than the 10 points commonly accepted as “very strong evidence” in favor of the model with the more negative value (Raftery 1995). Given that recent simulation studies with the models tested showed that differences in BIC values from -10 to 10 suggest that the alternative models fit “equally well” (Van Hulle et al. 2013), we chose to interpret the NL Main GxM model; however, the results from this model were interpreted with caution. As presented in Table 2 and Fig. 2b, the heritability of internalizing increased from 9% at the lowest level of affection to 32% at the highest level of affection. Shared environmental influences decreased from low ($c^2 = 34\%$) to high ($c^2 = 0\%$) levels of paternal affection. Nonshared environmental effects were weaker for low ($e^2 = 57\%$) levels of paternal affection, but increased at high levels of affection ($e^2 = 68\%$). This model also provide estimates for paternal affection: 36%

of the variance was due to genetic influences and 32 and 31% of the variance was due to shared and nonshared environmental influences, respectively.

Parenting behavior and alcohol use problems

The raw alcohol use problems score was not significantly correlated with raw scores for maternal discipline ($r = -.04$, $p = .20$), maternal affection ($r = .00$, $p = .98$), or paternal discipline ($r = .00$, $p = 1.00$). It was significantly correlated with paternal affection, though the correlation was small ($r = -.07$, $p = .01$).

For maternal discipline and maternal affection, the NL Main model provided the best fit to data according to AIC, BIC, and LRT (see Table 1). ACE estimates for these models were interpreted (see Table 2). For the NL Main model with maternal discipline, 38% of the variance in alcohol problems was accounted for by genetic influences, 0% of the variance was accounted for by shared environmental influences, and 62% of the variance was accounted for by nonshared environmental influences. The variance estimates for maternal discipline were identical to those found for the model with internalizing: 44% genetic, 11% shared environment, and 45% nonshared environment. For the NL Main model with maternal affection, genetic influences accounted for 37% of the variance in alcohol use problems and the remainder of the variance (63%) was accounted for by nonshared environmental influences; variance components for maternal affection were almost identical to those from the model with internalizing (43% genetic, 23% shared environment, 34% nonshared environment).

For paternal discipline and alcohol use, again, the NL Main model best fit the data. This model resulted in the lowest AIC (7110) and BIC (7155) and resulted in a nonsignificant loss of fit from the NL Main GxM model ($\chi^2_{diff} = 0.47$, 3 df, $p = .92$). The ACE estimates for the NL Main model indicate that 38% of the variance in alcohol problems is accounted for by genetic effects and the remainder (62%) of the variance is accounted for by nonshared environmental effects. Variance estimates for paternal discipline were identical to those found in the

Table 2 Estimates of unstandardized and standardized variance components for models

	SD	Variance components			Total variance	Proportions of variance		
		A	C	E		A (%)	C (%)	E (%)
<i>Internalizing models</i>								
Maternal discipline								
Maternal discipline		0.44	0.11	0.45	1.00	–	–	–
Internalizing		0.21	0.14	0.64	0.99	0.21	0.14	0.65
Maternal affection								
Maternal affection		0.44	0.23	0.34	1.01	–	–	–
Internalizing		0.14	0.18	0.64	0.96	0.15	0.18	0.67
Paternal discipline								
Paternal discipline		0.19	0.31	0.48	0.98	–	–	–
Internalizing	–2	0.02	0.50	0.43	0.95	0.02	0.53	0.45
	–1	0.10	0.19	0.50	0.79	0.13	0.24	0.63
	0	0.26	0.03	0.57	0.87	0.30	0.03	0.66
	1	0.50	0.01	0.65	1.17	0.43	0.01	0.56
	2	0.82	0.14	0.74	1.69	0.48	0.08	0.44
Paternal affection								
Paternal affection		0.36	0.32	0.31	0.99	–	–	–
Internalizing	–2	0.11	0.43	0.72	1.26	0.09	0.34	0.57
	–1	0.14	0.26	0.68	1.08	0.13	0.24	0.63
	0	0.18	0.13	0.64	0.94	0.19	0.14	0.67
	1	0.22	0.04	0.59	0.86	0.25	0.05	0.70
	2	0.26	0.00	0.56	0.82	0.32	0.00	0.68
<i>Alcohol use problem models</i>								
Maternal discipline								
Maternal discipline		0.44	0.11	0.45	1.00	–	–	–
Alcohol use problems		0.38	0.00	0.62	1.00	0.38	0.00	0.62
Maternal affection								
Maternal Affection		0.4	0.23	0.34	1.00	–	–	–
Alcohol use problems		0.37	0.00	0.62	0.99	0.37	0.00	0.63
Paternal discipline								
Paternal discipline		0.19	0.31	0.48	0.98	–	–	–
Alcohol use problems		0.38	0.00	0.62	1.00	0.38	0.00	0.62
Paternal affection								
Paternal affection		0.36	0.32	0.31	0.99	–	–	–
Alcohol use problems	–2	0.52	0.00	0.78	1.30	0.40	0.00	0.60
	–1	0.43	0.00	0.70	1.13	0.38	0.00	0.62
	0	0.36	0.00	0.62	0.98	0.36	0.00	0.64
	1	0.29	0.00	0.55	0.84	0.34	0.00	0.66
	2	0.22	0.00	0.48	0.71	0.32	0.00	0.68

A additive genetic, C shared environment, E nonshared environment

model with internalizing: 19 % genetic, 31 % shared environment, 48 % nonshared environment.

Finally, for paternal affection and alcohol problems, there was evidence suggesting that the NL Main GxM provided the best fit to data. It resulted in the lowest AIC (6950) and had a BIC value (7008) that was less than 10 points from the lowest BIC value (7000 for NL Main),

although going from the NL Main GxM to the NL Main model resulted in a nonsignificant loss of fit ($\chi^2_{\text{diff}} = 5.70, 3 \text{ df}, p = .13$). Therefore, as with paternal affection and internalizing, we chose to interpret the NL Main GxM model, but with caution. Genetic influences on alcohol problems decreased from low ($a^2 = 40\%$) to high ($a^2 = 32\%$) levels of paternal affection (see Table 2; Fig. 2c). Shared

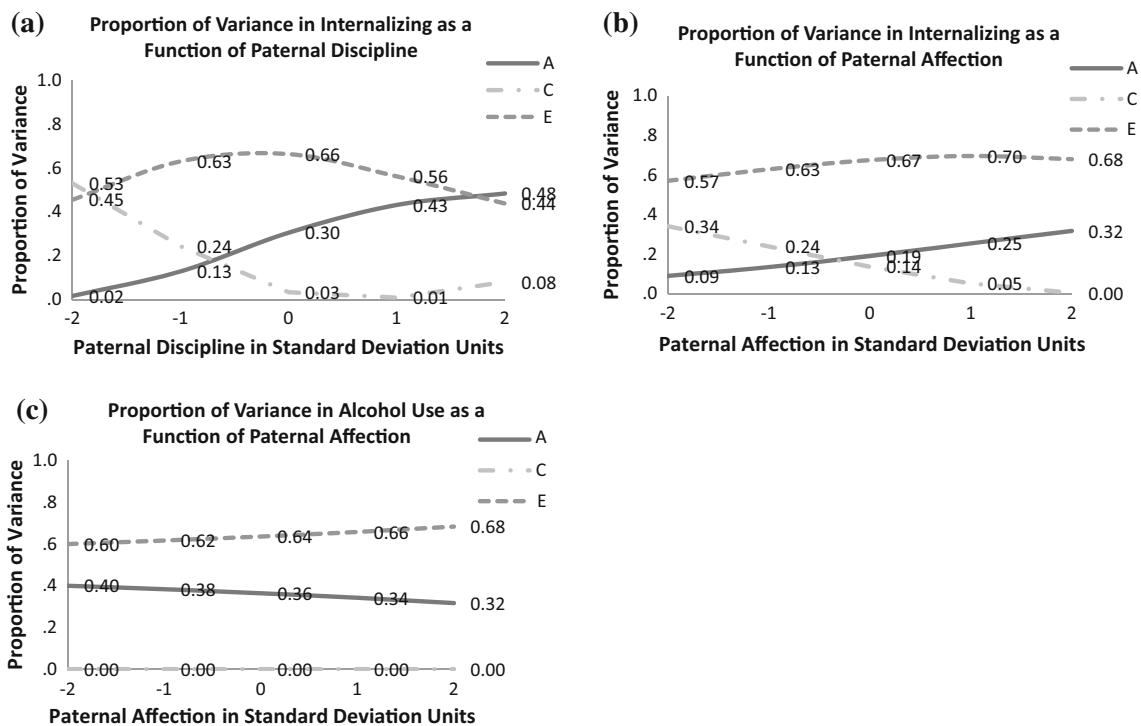


Fig. 2 **a** Proportion of variance in internalizing as a function of paternal discipline. **b** Proportion of variance in internalizing as a function of paternal affection. **c** Proportion of variance in alcohol use

problems as a function of paternal affection. *A* additive genetic, *C* shared environment, *E* nonshared environment

environmental influences were negligible across level of paternal affection. Unique environmental influences increased slightly from low ($e^2 = 60\%$) to high ($e^2 = 68\%$) levels of paternal affection. The estimates of genetic and environmental influences on paternal affection were identical to those found in the model with internalizing: 36% genetic, 32% shared environment, and 31% nonshared environment.

Discussion

Findings from a growing body of research have empirically established the presence of Gx E for major domains of psychopathology, including internalizing symptoms and alcohol use problems. One likely moderator of adult mental health that has not been fully explored to this point is the parent–child relationship. Parents who use strict discipline or are lacking in affection may put their genetically predisposed children at risk for mental health problems in adulthood. In the current study, we examined retrospective reports of parenting behavior—discipline and affection measured separately for mothers and fathers—as moderators of the genetic and environmental influences on internalizing symptoms and alcohol use problems in a nationally representative sample of adult twins. As has

been found previously (e.g., Rowe 1981), our results demonstrated that aspects of the parent–child relationship were moderately heritable. Our main findings indicated that father’s parenting behavior moderated the etiology of both internalizing symptoms and alcohol use problems. The genetic and environmental influences for internalizing symptoms and alcohol use problems differed depending on the level of affection received from a father during childhood, while genetic and environmental variance on internalizing symptoms also varied as a function of paternal discipline.

When paternal discipline was the moderator of the etiological influences on internalizing, there was a substantial increase in heritability (and genetic variance) from low to high levels of discipline. In line with other similar findings (Button et al. 2008; Lau and Eley 2008), our results support a diathesis–stress model, with greater genetic effects expressed in the “risky” environment—one where there is strict paternal discipline. To our knowledge, this is the first study to find evidence of paternal discipline moderating the genetic and environmental influences on internalizing psychopathology. Like extreme levels of maternal discipline (Lau and Eley 2008), paternal discipline may also have an effect on the etiology of internalizing symptoms by allowing for the expression of genetic influences.

Table 3 Parameter estimates and confidence intervals for best fitting models

	A_m	C_m	E_m	A_u	C_u	E_u	β_1	β_2	A_{ux}	C_{ux}	E_{ux}
<i>Internalizing</i>											
<i>Maternal discipline</i>											
NL main	0.66* (0.43–0.89)	-0.33 (-0.74 to 0.09)	0.67* (0.61–0.74)	-0.46* (-0.84 to -0.07)	0.37 (-0.03 to 0.77)	-0.80* (-0.87 to -0.73)	-0.01 (-0.08 to 0.07)	-0.04 (-0.09 to 0.01)			
<i>Maternal affection</i>											
NL Main	0.66* (0.48–0.85)	0.48* (0.25–0.72)	0.58* (0.52–0.63)	0.38 (-0.07 to 0.83)	0.42* (0.08–0.76)	0.80* (0.73–0.88)	0.13* (0.05–0.22)	-0.03 (-0.08 to 0.03)			
<i>Paternal discipline</i>											
NL main w/GxM	0.44* (0.12–0.77)	0.56* (0.34–0.78)	0.69* (0.63–0.76)	0.51* (0.32–0.71)	0.17 (-0.31 to 0.65)	0.76* (0.69–0.83)	-0.05 (-0.14 to 0.03)	-0.01 (-0.07 to 0.05)	0.20 (-0.06 to 0.45)	-0.27* (-0.49 to -0.05)	0.05 (-0.02 to 0.12)
<i>Paternal affection</i>											
NL main w/GxM	0.60* (0.41–0.79)	0.57* (0.38–0.76)	0.56* (0.50–0.61)	0.42* (0.00–0.84)	0.36 (-0.06–0.77)	0.80* (0.72–0.87)	0.14* (0.06–0.21)	0.02 (-0.04 to 0.08)	0.04 (-0.25 to 0.33)	-0.15 (-0.35 to 0.05)	-0.03 (-0.10 to 0.05)
<i>Alcohol use problems</i>											
<i>Maternal discipline</i>											
NL main	-0.66* (-0.89 to -0.43)	0.33 (-0.09 to 0.74)	-0.67* (-0.74 to -0.61)	-0.62* (-0.72 to -0.51)	0.00 (-1.36 to 1.36)	0.79* (0.71 to 0.86)	-0.01 (-0.09 to 0.07)	-0.01 (-0.06 to 0.04)			
<i>Maternal affection</i>											
NL main	-0.66* (-0.85 to -0.48)	-0.48* (-0.72 to -0.25)	-0.58* (-0.63 to -0.52)	-0.61* (-0.72 to -0.50)	0.00 (-1.32 to 1.32)	-0.79* (-0.86 to -0.71)	0.02 (-0.07 to 0.12)	0.02 (-0.04 to 0.08)			
<i>Paternal discipline</i>											
NL main	-0.44* (-0.76 to -0.12)	-0.56* (-0.78 to -0.34)	-0.69* (-0.76 to -0.63)	-0.62* (-0.72 to -0.51)	0.00 (-1.23 to 1.23)	-0.79* (-0.86 to -0.71)	-0.01 (-0.09 to 0.07)	0.01 (-0.04 to 0.07)			
<i>Paternal affection</i>											
NL main w/GxM	0.60* (0.41–0.79)	0.57* (0.38–0.76)	-0.56* (-0.61 to -0.50)	0.60* (0.49–0.71)	0.00 (-1.09 to 1.09)	0.80* (0.72–0.86)	0.08* (0.00–0.16)	0.06 (0.00–0.13)	-0.06 (-0.17 to 0.05)	0.00 (-0.22 to 0.22)	-0.05 (-0.12 to 0.02)

A_m genetic path estimate for parent-child relationship quality, C_m shared environmental path estimate for parent-child relationship quality, E_m nonshared environmental path estimate for parent-child relationship quality, A_u genetic path estimate unique to psychopathology, C_u shared environmental path estimate unique to psychopathology, E_u nonshared environmental path estimate unique to psychopathology, β_1 main effect of parent child-relationship quality on psychopathology, β_2 main effect squared of parent child-relationship quality on psychopathology, A_{ux} moderator of genetic path to psychopathology, C_{ux} moderator of common environmental path to psychopathology, E_{ux} moderator of nonshared environmental path to psychopathology

* $p < 0.05$

When paternal affection was the moderator of genetic and environmental influences on internalizing, the genetic variance (and heritability) increased from low to high levels of affection. Given the mixed evidence in support of biometric moderation, it is recommended that these results be interpreted with caution. Indeed, a similar study failed to find a moderating effect of a summary parental affection score on the etiological influences on depressive symptoms (Feinberg et al. 2007). If replicated, the current findings may be suggestive of a social push model, with greater genetic influences in the most “enriched” environment (i.e., one marked by the highest levels of affection). Individuals who are immersed in an advantageous environment, such as an environment with high paternal affection, may exhibit internalizing symptoms despite a warm and caring father-child relationship because they are unable to counteract a genetic predisposition.

Of note, the proportion of variance in internalizing due to shared environmental influences was highest in the least affectionate and least disciplined environments. Any non-zero estimates of the shared environment on adult symptoms of psychopathology are notoriously difficult to find; thus, the fact that the proportion of variance in internalizing due to the shared environment was estimated at 53 % (from a model with paternal discipline) and at 34 % (from a model with paternal affection) at extremely low levels of the moderator is important. These substantial estimates can be interpreted as meaning that the type of father parenting experienced by both twins had a moderate to strong impact on whether each sibling later reported internalizing symptoms. This suggests that *lack of father engagement per se*, even when that engagement is controlling and strict, has an important role in the development of mental illness.

Paternal affection also significantly moderated the etiological influences on alcohol use problems. The drop in genetic variance from low to high levels of affection was paralleled by a concurrent drop in nonshared environmental variance, such that there was greater overall variance in alcohol use problems when paternal affection was lower. Our findings provide evidence for a diathesis-stress model of etiology, wherein genetic influences emerge in an environment where the father is less warm and loving. This is consistent with a majority of studies investigating biometric moderation of alcohol use problems and related phenotypes, which tend to find support for a diathesis-stress model of etiology (Young-Wolff et al. 2011).

No moderation of internalizing or alcohol use problems as a function of maternal affection or maternal discipline was found. Hicks et al. (2009a) reported moderation of the ACE components of internalizing as a function of mother-child relationship problems, but only for the nonshared environmental parameter. Differences could be due to

several factors, including different measures of the parent-child relationship, different mean ages of the samples when assessed for psychopathological symptoms, or the cross-sectional nature of that study compared to the retrospective measurement of parent-child relationship utilized in the current study.

Even though this study failed to find moderation of the etiological influences on internalizing and alcohol use problems as a function of the mother-child relationship, this does not suggest that only the paternal relationship is important in the development of mental health problems. Our findings will certainly need to be replicated in other samples. The mother-child relationship may contribute to later psychopathological symptoms but may not have (or maintain) a moderating effect on the etiological influences in psychopathology when one is well into adulthood. Our findings are, however, consistent with research supporting the importance of the father in the home for predicting later offspring psychopathology (Blazei et al. 2008; Harold et al. 2012).

This study is not without limitations. Foremost, our measure of the parent-child relationship was ascertained during adulthood in our sample, which ranged in age from 25 to 74 at time of data collection. Thus, symptoms of adult mental illness may have colored participants' recollection of the parent-child relationship and this bias may have been more similar between MZ twin pairs, compared to DZ pairs. If this is the case, estimates for genetic variance may be inflated and estimates of the nonshared environment may actually contain a large degree of measurement error. It is also possible that individuals from different generations might have had different interpretations of what “discipline” and “affection” entailed in regards to their relationship with their parents. Although there are limitations surrounding use of this measure, evidence suggests that adult retrospective reports of parenting behavior are consistent with childhood assessments (Brewin et al. 1993; Wilhelm et al. 2005), and retrospective reports of parenting predicts well-being across the life-course (Russek and Schwartz 1997). Nevertheless, future studies should endeavor to replicate these findings using prospective studies with more homogenous and recently studied samples. Second, our measures of mental health problems were all self-report and different results might have been obtained using clinician- or informant-report measures. Third, the size of the sample used in the current analyses was somewhat small compared to other twin studies that use biometric moderation models. Despite this relatively small sample size there was still enough power to find moderation of paternal relationship quality on the etiological influences on adult mental health. It is unknown if maternal relationship quality would have acted as a moderator if there was a larger sample size, however.

In summary, our results reveal the importance of the father-child relationship for understanding the etiology of symptoms of psychopathology later in life. Following a diathesis-stress model of psychopathology, greater genetic influences on internalizing were found at extremely high levels of paternal discipline. A diathesis-stress model was also supported for the effect of paternal affection on alcohol use problems. The effect of paternal affection on the etiological components of internalizing, however, was best explained by a social push model. Our findings continue to add to what is known about the gene-environment interplay between aspects of the parent-child relationship and later psychopathology. Affection and discipline, however, while important constructs, are only two dimensions of parenting; future work is needed to determine whether and in what way other aspects of the parent-child relationship may impact the etiology of the child's future mental health.

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Conflict of Interest Susan C. South, Amber M. Jarnecke declare that they have no conflict of interest.

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